# Gov 50: 15. Multiple Regression and Interpretation 

Matthew Blackwell

Harvard University

## Roadmap

1. Multiple regression
2. Categorical independent variables

1/ Multiple regression

## Multiple predictors

What if we want to predict $Y$ as a function of many variables?

$$
\text { seat_change }_{i}=\alpha+\beta_{1} \text { approval }_{i}+\beta_{2} \text { rdi_change }_{i}+\epsilon_{i}
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- Statistical control in a cross-sectional study.


## Multiple regression in $\mathbf{R}$

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## Call:
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## Coefficients:
## (Intercept) approval rdi_change
## -117.23 1.53 3.22
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$\begin{array}{llll}\text { \#\# } & \text {-117.23 } & 1.53 & 3.22\end{array}$

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- $\hat{\beta}_{1}=1.53$ : average increase in seat change for additional percentage point of approval, holding RDI change fixed
- $\hat{\beta}_{2}=3.217:$ average increase in seat change for each additional percentage point increase of RDI, holding approval fixed


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## Least squares with multiple regression

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- The same exact way as before: minimize prediction error!
- Residuals (aka prediction error) with multiple predictors:

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Y_{i}-\widehat{Y}_{i}=\text { seat_change }_{i}-\hat{\alpha}-\hat{\beta}_{1} \text { approval }_{i}-\hat{\beta}_{2} \text { rdi_change }_{i}
$$

- Find the coefficients that minimizes the sum of the squared residuals:

$$
\mathrm{SSR}=\sum_{i=1}^{n} \hat{\epsilon}_{i}^{2}=\left(Y_{i}-\hat{\alpha}-\hat{\beta}_{1} X_{i 1}-\hat{\beta}_{2} X_{i 2}\right)^{2}
$$

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## Model fit with multiple predictors

- $R^{2}$ mechanically increases when you add a variables to the regression.
- But this could be overfitting!!
- Solution: penalize regression models with more variables.
- Occam's razor: simpler models are preferred
- Adjusted $R^{2}$ : lowers regular $R^{2}$ for each additional covariate.
- If the added covariates doesn't help predict, adjusted $R^{2}$ goes down!


## Comparing model fits

```
library(broom)
fit.app <- lm(seat_change ~ approval, data = midterms)
glance(fit.app) |>
    select(r.squared, adj.r.squared, sigma)
```

| \#\# \# A tibble: $1 \times 3$ |  |  |  |
| :--- | ---: | ---: | ---: |
| \#\# | r.squared adj.r.squared | sigma |  |
| \#\# | <dbl> | <dbl> | <dbl> |
| \#\# 1 | 0.450 | 0.418 | 16.9 |

```
glance(mult.fit) |>
    select(r.squared, adj.r.squared, sigma)
```

| \#\# \# A tibble: $1 \times 3$ |  |  |  |
| :--- | ---: | ---: | ---: |
| \#\# | r.squared adj.r.squared | sigma |  |
| \#\# | <dbl> | <dbl> | <dbl> |
| \#\# 1 | 0.468 | 0.397 | 16.7 |

## Predicted values from $\mathbf{R}$

We could plug in values into the equation, but R can do this for us. The \{modelr\} package gives some functions that allow us to predictions in a tidy way:

Let's use add_predictions( ) to predict the 2022 results

```
library(modelr)
midterms |>
    filter(year == 2022) |>
    add_predictions(mult.fit)
```

\#\# \# A tibble: $1 \times 7$
\#\# year president party approval seat_change rdi_cha~1 pred
\#\# <dbl> <chr> <chr> <dbl> <dbl> <dbl> <dbl>
\#\# 12022 Biden D 42 NA -0.003 -53.2
\#\# \# ... with abbreviated variable name 1: rdi_change

## Predictions from several models

The gather_predictions( ) will return one row for each model passed to it with the prediction for that model:

```
midterms |>
    filter(year == 2022) |>
    gather_predictions(fit.app, mult.fit)
```

\#\# \# A tibble: 2 x 8
\#\# model year presi~1 party appro~2 seat_~3 rdi_c~4 pred
\#\# <chr> <dbl> <chr> <chr> <dbl> <dbl> <dbl> <dbl>
\#\# 1 fit.app 2022 Biden $\quad 42$ NA -0.003 -36.9
\#\# 2 mult.fit 2022 Biden 42 NA -0.003 -53.2
\#\# \# ... with abbreviated variable names 1: president,
\#\# \# 2: approval, 3: seat_change, 4: rdi_change

## Predictions from new data

What about predicted values not in data?

```
tibble(approval = c(50, 75), rdi_change = 0) |>
    gather_predictions(fit.app, mult.fit)
```

\#\# \# A tibble: $4 \times 4$
\#\# model approval rdi_change pred
\#\# <chr> <dbl> <dbl> <dbl>
\#\# 1 fit.app $50 \quad 0$-25.6
\#\# 2 fit.app 7509.92
\#\# 3 mult.fit $50 \quad 0-40.9$
$\begin{array}{llll}\text { \#\# } 4 \text { mult.fit } 75 & 0 & -2.79\end{array}$

## Predictions from augment ()

We can also get predicted values from the augment ( ) function using the newdata argument:

```
newdata <- tibble(approval = c(50, 75), rdi_change = 0)
augment(mult.fit, newdata = newdata)
```

\#\# \# A tibble: $2 \times 3$
\#\# approval rdi_change .fitted
\#\# <dbl> <dbl> <dbl>
\#\# $1000-40.9$
$\begin{array}{llll}\# \# 2 & 75 & 0 & -2.79\end{array}$

2/ Categorical independent variables

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- Program had support from most parties.
- Was implemented in a nonpartisan fashion.


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- Welfare \$\$ given if kids enrolled in schools, get regular check-ups, etc.
- Do these programs have political effects?
- Program had support from most parties.
- Was implemented in a nonpartisan fashion.
- Would the incumbent presidential party be rewarded?


## The data

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| treatment | early Progresa (1) or late Progresa (0) <br> pri2000s |
| PRI votes in the 2000 election as a share of adults <br> in precinct |  |
| t2000 | turnout in the 2000 election as share of adults in <br> precinct |

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```
library(qss)
data("progresa", package = "qss")
cct <- as_tibble(progresa) |>
    select(treatment, pri2000s, t2000)
cct
```



## Difference in means estimates

## Does CCT affect turnout?

```
cct |> group_by(treatment) |>
    summarize(t2000 = mean(t2000)) |>
    pivot_wider(names_from = treatment, values_from = t2000) |>
    mutate(ATE = `1` - `0`)
```

```
## # A tibble: 1 x 3
## `0` `1` ATE
## <dbl> <dbl> <dbl>
## 1 63.8 68.1 4.27
```


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- Slope $\hat{\beta}$ is the difference-in-means of $Y$ between $X=1$ group and $X=0$ group.

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\hat{\beta}=\bar{Y}_{\text {treated }}-\bar{Y}_{\text {control }}
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$$
\hat{\beta}=\bar{Y}_{\text {treated }}-\bar{Y}_{\text {control }}
$$

- If there are other independent variables, this becomes the difference-in-means controlling for those covariates.


## Linear regression for experiments

- Under randomization, we can estimate the ATE with regression:

```
cct |> group_by(treatment) |>
    summarize(pri2000s = mean(pri2000s)) |>
    pivot_wider(names_from = treatment, values_from = pri2000s) |>
    mutate(ATE = `1` - `0`)
```

| \#\# \# A tibble: | 1 | x | 3 |  |
| :--- | :--- | :--- | :--- | :--- |
| \#\# | `0` | `1. | ATE |  |
| \#\# | <dbl> | <dbl> | <dbl> |  |
| \#\# | 1 | 34.5 | 38.1 | 3.62 |

lm(pri2000s ~ treatment, data $=c c t) \mid>\operatorname{coef}()$

| \#\# (Intercept) | treatment |  |
| :--- | ---: | ---: |
| \#\# | 34.49 | 3.62 |

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| Unit | Party | Democrat | Republican | Independent |
| :--- | :--- | :--- | :--- | :--- |
| 1 | Democrat | 1 | 0 | 0 |
| 2 | Democrat | 1 | 0 | 0 |
| 3 | Independent | 0 | 0 | 1 |
| 4 | Republican | 0 | 1 | 0 |
| $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ |

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| 4 | Republican | 0 | 1 | 0 |
| $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ | $\vdots$ |

- Then include all but one of these binary variables:

$$
\text { turnout }_{i}=\alpha+\beta_{1} \text { Republican }_{i}+\beta_{2} \text { Independent }_{i}+\varepsilon_{i}
$$

## Interpreting categorical variables

turnout $_{i}=\alpha+\beta_{1}$ Republican $_{i}+\beta_{2}$ Independent $_{i}+\varepsilon_{i}$

- $\hat{\alpha}$ : average outcome in the omitted group/baseline (Democrats).


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- $\hat{\beta}$ coefficients: average difference between each group and the baseline.


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- $\hat{\beta}_{1}$ : average difference in turnout between Republicans and Democrats


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- $\hat{\beta}$ coefficients: average difference between each group and the baseline.
- $\hat{\beta}_{1}$ : average difference in turnout between Republicans and Democrats
- $\hat{\beta}_{2}$ : average difference in turnout between Independents and Democrats


## CCES data

## library(gov50data) <br> cces_2020

```
## # A tibble: 51,551 x 6
## gender race educ pid3 turno~1 pres_~2
## <fct> <fct> <fct>
## 1 Male White 2-year
## 2 Female White Post-grad
## 3 Female White 4-year
## 4 Female White 4-year
## 5 Male White 4-year
## 6 Male White Some college
## 7 Male Black Some college
## 8 Female White Some college
    <fct> <dbl> <fct>
    Republ~ 1 Donald~
    Democr~ NA <NA>
    Indepe~ 1 Joe Bi~
    Democr~ 1 Joe Bi~
    Indepe~ 1 Other
    Republ~ 1 Donald~
    Not su~ NA <NA>
    Indepe~ 1 Donald~
## 9 Female White High school graduate Republ~ 1 Donald~
## 10 Female White 4-year Democr~ 1 Joe Bi~
## # ... with 51,541 more rows, and abbreviated variable names
## # 1: turnout_self, 2: pres_vote
```


## Categorical variables in the CCES data

```
turnout_pred <- lm(turnout_self ~ pid3, data = cces_2020)
turnout_pred
```

\#\#
\#\# Call:
\#\# lm(formula $=$ turnout_self ~ pid3, data = cces_2020)
\#\#
\#\# Coefficients:

| \#\# | (Intercept) | pid3Republican | pid3Independent |
| :--- | ---: | ---: | ---: |
| \#\# | 0.9635 | -0.0103 | -0.0394 |
| \#\# | pid3Other | pid3Not sure |  |
| \#\# | -0.0066 | -0.3331 |  |

## What $R$ does internally with factor variables in lm

```
cces_2020 |> drop_na(turnout_self, pid3) |> select(pid3) |> pull() |>
    head()
```

\#\# [1] Republican Independent Democrat Independent
\#\# [5] Republican Independent
\#\# 7 Levels: Democrat Republican Independent ... not asked
model.matrix(turnout_pred) |>
head()

| \#\# | (Intercept) | pid3Republican | pid3Independent | pid30ther |
| :--- | ---: | ---: | ---: | ---: |
| \#\# | 1 | 1 | 1 | 0 |
| \#\# | 3 | 1 | 0 | 1 |
| \#\# | 4 | 1 | 0 | 0 |
| \#\# | 5 | 1 | 0 | 1 |


| \#\# | pid3Not sure |
| :--- | ---: |
| \#\# 1 | 0 |
| \#\# 3 | 0 |
| \#\# 4 | 0 |
| \#\# 5 | 0 |
| \#\# 6 | 0 |

